

# A Online appendix (Not for publication)

## A.1 Variable definition and sources

**Table A-1: Variable definition and sources**

Variable	Description
<i>Age</i>	Respondent's age in years.
<i>Agree with bribery</i>	Equals 1 if respondent "totally agrees" or "agrees" with the statement: "As things are, sometimes paying a bribe is justified."
<i>Anti-Corruption Index</i>	Summary measure for the performance of strategic anti-corruption standards in municipalities using measures of: internal control, recruitment, administrative management systems and accountability. Computed by the <i>General Attorney</i> , available as <i>Índice de Gobierno Abierto</i> .
<i>Authorities violate the law</i>	Equals 1 if respondent "totally agrees" or "agrees" with the statement: "To capture criminals, authorities should sometimes violate the law."
<i>Catholic</i>	Equals 1 for Catholics (self-declared).
<i>Contributes to social security</i>	Equals 1 if respondent pays contributions to the social security system (thus excluding members in the subsidized social security system).
<i>Dark</i>	Respondent's skin color based on color palette (assessed by interviewer). The palette is numbered from 1 to 11 (1 = lightest color). <i>Dark</i> Equals 1 for colors greater than or equal to 5.
<i>Education</i>	Educational attainment. Equals 1 if respondent has some secondary education or more.
<i>Employed household head</i>	Equals 1 if household head was employed in the week preceding the survey.
<i>Employment</i>	Equals 1 if respondent was employed the week previous to the survey.
<i>Evangelical/Pentecostal</i>	Equals 1 for Evangelical/Pentecostal (self-declared).
<i>FEA</i>	Equals 1 if household is a <i>Familias en acción</i> (main conditional cash transfer for the poor with school-age children) beneficiary.
<i>Formal credit</i>	Equals 1 if household has any formal credit.
<i>Fractionalization</i>	$F_j = 1 - \sum_{i=1}^N \pi_{ij}^2$ , where $\pi_{ij}$ is the vote share for the mayoral candidate (in 2011) $i$ in municipality $j$ . See <a href="#">Montalvo and Reynal-Querol (2005)</a> . <a href="#">Pachón and Sánchez (2014)</a> .

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**Table A-1 – Variable definition and sources, continued from previous page**

<b>Variable</b>	<b>Description</b>
<i>Frequency of voting</i>	Equals 1 if respondent “Always votes in elections” or “Votes in most elections”, 0 otherwise (“Rarely votes in elections” or “Never voted”).
<i>Get help</i>	Equals 1 if household answers “yes” to at least one of the following questions: “During the past 12 months, do any members of the household receive money or in-kind aid...” a. “...from relatives or friends living in Colombia?”, b. “from relatives or friends living abroad?” c. “for alimony?”, d. “from international organizations (WFP, UNICEF, ICRC)?”, e. “from NGOs?”, f. “from the church or other religious organizations?”, g. “from other persons, entities or organizations?”.
<i>Gini of land properties</i>	Municipal land Gini. <i>Cede Panel</i> .
<i>Gov. against inequality</i>	Equals 1 if respondent “totally agrees” or “agrees” with the statement: “The government should implement strong policies to reduce inequality between rich and poor.”
<i>Government role</i>	Equals 1 if respondent “totally agrees” or “agrees” with the statement: “Government is primarily responsible for ensuring the welfare of the people.”
<i>Guerrillas</i>	Number of violent events per year perpetrated by guerrillas per 100,000 inhabitants (average 2010–2012). <i>Conflict analysis resource center (CERAC) and Universidad del Rosario</i> .
<i>HH expenses</i>	Per capita household expenses. See <a href="#">Bernal et al. (2014)</a> .
<i>HH food expenses</i>	Household expenses in food (Colombian pesos). See <a href="#">Bernal et al. (2014)</a> .
<i>Homeowner</i>	Equals 1 if the household residence is “own, fully paid” or “own, being paid”. Equals 0 otherwise (“rented” or “in usufruct or other type of tenure”).
<i>Homicide rate</i>	Homicide rate by 100,000 inhabitants (average 2010–2012). <i>Medicina Legal</i> .
<i>Household with spouse</i>	Equals 1 if household is inhabited by household head and spouse.
<i>Independent</i>	Equals 1 if working independently is the most important job during the previous month.
<i>Justice into own hands</i>	Equals 1 if respondent “totally agrees” or “agrees” with the statement: “When the government does not punish criminals, it is okay that people take justice into their own hands.”
<i>Lands</i>	Equals 1 if respondent reports owning land.

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**Table A-1 – Variable definition and sources, continued from previous page**

<b>Variable</b>	<b>Description</b>
<i>Left and Right Ideology</i>	Equals 1 if respondent reports “left” (“right”) or “center-left” (“center-right”) when asked: “Often, people speak of political leanings from left and right. According to the meaning that the terms ‘left’ and ‘right’ have for you, which political tendency do you sympathize with?”
<i>Left and Right dominated</i>	Equals 1 if municipal mayor belongs to a left- (right-) leaning party. Parties are coded following <a href="#">Fergusson, Querubín, et al. (2017)</a> . <a href="#">Pachón and Sánchez (2014)</a> .
<i>Male household head</i>	Equals 1 if household head is male.
<i>Neighbor cell phones</i>	Equals 1 if person has the cell phone number of at least half of her neighbors.
<i>Neighbor loans</i>	Equals 1 if a person thinks that at least half of her neighbors would lend her money.
<i>No debts</i>	Equals 1 if household has not any type of debt. Equals 0 otherwise (“debt without credit”, “formal credit”, “informal credit” or “formal-informal credit”).
<i>No sewage</i>	Equals 1 if household dwelling has no sewage system.
<i>Non left-right contender</i>	Equals 1 if municipality had no right-wing or left-wing party (among the top two) in the most recent mayoral election (classification of <a href="#">Fergusson, Querubín, et al. (2017)</a> ).
<i>Not in organization</i>	Equals 1 if respondent does not belong to any organization (options included are <i>Juntas de acción comunal</i> , charity organization, community organization, religious organization, organizations supported or promoted by the state, ethnic organization, educational organization, labor union, cooperative of work or union of producers, organization of environment conservation, cultural or sports organization, other).
<i>Negative reciprocity</i>	Equals 1 if respondent “totally agrees” or “agrees” with the statement: “Whoever hurts me, pays for it.”
<i>Nuclear family</i>	Equals 1 household is comprised of: household head and spouse, with or without children; or, household head without spouse but with children).
<i>Other religion</i>	Equals 1 for believers of religions other than Catholic, Evangelical, or Pentecostal (self-declared).
<i>Overcrowded</i>	Equals 1 if ratio of number of residents to number of bedrooms is greater than three in rural households, or greater than or equal to three in urban households.

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**Table A-1 – Variable definition and sources, continued from previous page**

<b>Variable</b>	<b>Description</b>
<i>Own welfare</i>	Equals 1 if respondent “totally agrees” or “agrees” with the statement: “Each individual is responsible for their own welfare.”
<i>Paramilitaries</i>	Number of violent events per year perpetrated by paramilitaries per 100,000 inhabitants (average 2010–2012). <i>Conflict analysis resource center (CERAC) and Universidad del Rosario.</i>
<i>Party identity</i>	Equals 1 if respondent answers “yes” to: “At the moment, do you sympathize with any political party?”
<i>Party recall</i>	Equals 1 if respondent answers “yes” to: “Do you remember which party you vote for mayor of your city or municipality?”
<i>People in household</i>	Number of household residents.
<i>Persuasion</i>	Equals 1 if respondent answers “frequently,” “sometimes”, or “rarely” (i.e., leaves out “never”) to “During elections, some people try to convince others to vote for a particular party or candidate. How often have you tried to convince others to vote for a party or candidate?”
<i>Polarization</i>	<b>Reynal-Querol (2002)</b> polarization index. $P_j = 1 - \sum_{i=1}^N \pi_{ij} \left( \frac{1/2 - \pi_{ij}}{1/2} \right)^2$ , where $\pi_{ij}$ is the vote share for the mayoral candidate (in 2011) $i$ in municipality $j$ . <i>Pachón and Sánchez (2014).</i>
<i>Pop. density</i>	Population divided by total area ( $km^2$ ) in the municipality.
<i>Popular vote</i>	Equals 1 if respondent “totally agrees” or “agrees” with the statement: “It is important that rulers are elected by popular vote.”
<i>Positive reciprocity</i>	Equals 1 if respondent “totally agrees” or “agrees” with the statement: “You always have to help those who help you.”
<i>Public Transparency</i>	Summary measure of public transparency (lower levels in this indicator imply a higher risk of corruption) based on three criteria: visibility (management of public information, open data and access to information on the website), transparency in public procurement, and local governments capacity to apply sanctions and compliance delivering information to control and regulation agencies. Computed by <i>Transparencia por Colombia</i> , available as <i>Índice de Transparencia Departamental</i> .
<i>Regions</i>	Regions included in fixed effects. Urban regions in the survey are: <i>Atlántica, Oriental, Central, Pacífica, Bogotá</i> . Rural regions include: <i>Atlántica-Media, Cundi-Boyacense, Eje Cafetero, Centro-Oriente</i> .

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**Table A-1 – Variable definition and sources, continued from previous page**

<b>Variable</b>	<b>Description</b>
Risk of Electoral Fraud	Risk of fraud based on atypical electoral participation, limitations on electoral competition, and atypical levels of null votes and unmarked ballots. Classifies municipalities in four categories: “No Risk”, “Medium risk”, “High Risk” and “Extreme Risk”. In the figure, we aggregate the former two categories as “Low Risk” and the latter two categories as “High Risk”. Index computed by The Colombian Electoral Observation Mission (Misión de Observación Electoral, MOE).
<i>Rural population</i>	Proportion of rural population in the municipality (average 2006–2008). DANE.
<i>Savings</i>	Equals 1 if respondent answers “yes” to: “Do you usually save some of the income you receive?”.
<i>Secret ballot</i>	Equals 1 if respondent answers “yes” to: “Do you think that the ballot is secret?”
<i>Send help</i>	Equals 1 if household answers “yes” to at least one of the following questions: “During the past 12 months, did any members of the household send money or in-kind aid...” a. “to relatives or friends who live in Colombia?”, b. “to relatives or friends who live abroad?”, c. “for alimony?” d. “to other persons, entities or organizations?”.
<i>Shock</i>	Equals 1 if household reports any major destabilizing negative event during the previous three years.
<i>Social program beneficiary</i>	Equals 1 if household benefits from any of the following programs: <i>Familias en acción</i> (main conditional cash transfer for the poor with school-age children), programs for the elderly, <i>SENA</i> training programs, <i>Red Juntos - Unidos</i> (program that provides social services to displaced families with the lowest levels of poverty), <i>ICBF</i> programs for children, aid for displaced people, support to households affected by natural disasters, or “other programs”.
<i>Social security</i>	Equals 1 if respondent is affiliated to social security.
<i>State presence</i>	Raw total of local state agencies, local municipality employees, and national-level municipality employees (per capita in 1995). <i>Acemoglu, Garcia-Jimeno, and Robinson (2015)</i> .
<i>Stratum 1, 2</i>	Socio-economic stratum, based on classification of household residence (used to target utility subsidies).
<i>Use of violence</i>	Equals 1 if respondent “totally agrees” or “agrees” with the statement: “Sometimes the use of violence is justified.”

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**Table A-1 – Variable definition and sources, continued from previous page**

<b>Variable</b>	<b>Description</b>
<i>Vote for the same party</i>	Equals 1 if respondent “Always votes for the same party” or “Almost always votes for the same party”. Equals 0 otherwise (“Votes for different parties” or “Always votes blank”).
Voted in the last election	Equals 1 if respondent voted in the last election for mayors (in 2011), 0 otherwise.
<i>Wealth</i>	First principal component following a <i>principal component analysis</i> on a set of reported household assets and dwelling characteristics. See <a href="#">Bernal et al. (2014)</a> .
<i>Win margin</i>	Difference between the vote shares of the winner and runner-up in the 2011 mayoral election. <a href="#">Pachón and Sánchez (2014)</a> .
<i>Woman</i>	Equals 1 if respondent is female.

**Notes:** Source is Elca 2013 unless otherwise stated at the end of each description.

## A.2 Balance between treatment and control groups in list experiments

We corroborate that respondents assigned to treatment and control lists, and the direct question, have similar observable characteristics. For a set of observables  $\mathbf{X}$ , we check both the bivariate relationship between group assignment and observables:

$$\Pr(T_i = m) = f(x_i\beta_m) \quad \text{with } x_i \in \mathbf{x},$$

and the multivariate regression,

$$\Pr(T_i = m) = f(\mathbf{x}'\beta_m),$$

where  $m$  represents each group (*Treatment*, *Control 1*, and *Control 2*). We estimate the marginal effects of multinomial probit models.

Since randomization was stratified at the regional level, in both types of regressions we include region fixed effects. We also estimated separate regressions for each region, with similar results, but present only these aggregate results to save space. Similarly, we also estimated simple probit and linear probability models for dichotomous indicators of each treatment condition as the dependent variable, and again found no systematic evidence of imbalance.

Table [A-2](#) shows balance using observables in 2010 and Table [A-3](#) in 2013 for the clientelism experiment.

**Table A-2: Balance on covariates at the baseline (2010): vote-buying list experiment**

Variables	Urban sample						Rural sample					
	Bivariate			Multivariate			Bivariate			Multivariate		
	Treatment	Control 1	Control 2	Treatment	Control 1	Control 2	Treatment	Control 1	Control 2	Treatment	Control 1	Control 2
Age	0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	0.001 (0.00)	-0.001** (0.00)	0.001 (0.00)	0.001 (0.00)	-0.001** (0.00)	0.000 (0.00)
Male household head	-0.012 (0.02)	-0.008 (0.02)	0.021 (0.02)	0.012 (0.03)	-0.024 (0.02)	0.012 (0.03)	0.002 (0.02)	-0.012 (0.02)	0.010 (0.02)	0.008 (0.03)	-0.020 (0.03)	0.012 (0.03)
Education	-0.001 (0.00)	0.001 (0.00)	-0.000 (0.00)	0.000 (0.00)	0.002 (0.00)	-0.002 (0.00)	-0.002 (0.00)	0.002 (0.00)	-0.000 (0.00)	-0.000 (0.00)	-0.001 (0.00)	0.001 (0.00)
Employed household head	-0.021 (0.02)	0.010 (0.02)	0.011 (0.02)	-0.018 (0.02)	0.016 (0.02)	0.003 (0.02)	-0.001 (0.02)	0.019 (0.02)	-0.017 (0.02)	-0.000 (0.02)	0.017 (0.02)	-0.016 (0.02)
Savings	-0.009 (0.02)	-0.021 (0.02)	0.030 (0.02)	-0.008 (0.02)	-0.020 (0.02)	0.028 (0.02)	0.012 (0.03)	0.016 (0.02)	-0.028 (0.03)	0.021 (0.03)	0.008 (0.02)	-0.029 (0.03)
Not in organization	0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	-0.001 (0.00)	-0.001 (0.00)	0.002* (0.00)	-0.001 (0.00)	-0.001 (0.00)	0.002 (0.00)
Social security	0.006 (0.03)	0.013 (0.03)	-0.019 (0.03)	0.014 (0.03)	0.008 (0.03)	-0.022 (0.03)	0.009 (0.04)	-0.002 (0.04)	-0.007 (0.03)	0.014 (0.04)	-0.008 (0.04)	-0.006 (0.03)
Contributes to social security	-0.028* (0.01)	0.007 (0.01)	0.021 (0.01)	-0.033* (0.02)	0.011 (0.02)	0.021 (0.02)	-0.011 (0.02)	-0.015 (0.02)	0.026 (0.02)	-0.002 (0.03)	-0.028 (0.03)	0.030 (0.03)
Household with spouse	0.022 (0.01)	0.000 (0.02)	-0.022 (0.02)	0.028 (0.03)	-0.023 (0.02)	-0.006 (0.02)	0.006 (0.02)	0.002 (0.02)	-0.008 (0.02)	0.013 (0.03)	0.004 (0.03)	-0.018 (0.02)
Wealth	-0.002 (0.01)	-0.002 (0.01)	0.004 (0.01)	0.006 (0.01)	-0.008 (0.01)	0.002 (0.01)	-0.017** (0.01)	0.008 (0.01)	0.009 (0.01)	-0.013 (0.01)	0.002 (0.01)	0.011 (0.01)
People in household	0.001 (0.00)	0.004 (0.00)	-0.005 (0.00)	0.001 (0.00)	0.001 (0.00)	-0.003 (0.01)	0.005 (0.00)	-0.000 (0.00)	-0.004 (0.00)	0.002 (0.01)	0.003 (0.01)	-0.005 (0.01)
Overcrowded	-0.013 (0.02)	0.020 (0.02)	-0.007 (0.02)	-0.016 (0.02)	0.032 (0.02)	-0.016 (0.02)	-0.020 (0.02)	-0.010 (0.02)	0.030 (0.02)	-0.018 (0.02)	-0.004 (0.02)	0.023 (0.02)
Homeowner	0.003 (0.01)	-0.012 (0.01)	0.009 (0.01)	0.003 (0.02)	-0.020 (0.01)	0.017 (0.01)	-0.008 (0.01)	-0.002 (0.02)	0.009 (0.01)	-0.012 (0.02)	0.004 (0.02)	0.008 (0.02)
No debts	-0.013 (0.01)	-0.005 (0.01)	0.018 (0.01)	-0.009 (0.02)	-0.001 (0.02)	0.010 (0.02)	-0.007 (0.02)	0.033** (0.01)	-0.026* (0.01)	-0.005 (0.02)	0.020 (0.02)	-0.014 (0.02)
Shock	0.001 (0.02)	0.020 (0.01)	-0.020 (0.02)	0.003 (0.02)	0.019 (0.02)	-0.023 (0.01)	-0.015 (0.01)	-0.004 (0.01)	0.019 (0.02)	-0.015 (0.02)	-0.005 (0.01)	0.020 (0.02)
HH expenses (log)	-0.004 (0.01)	0.005 (0.01)	-0.001 (0.01)	0.005 (0.02)	0.008 (0.02)	-0.014 (0.02)	-0.003 (0.01)	0.010 (0.01)	-0.007 (0.01)	-0.015 (0.01)	0.030 (0.02)	-0.014 (0.02)
HH food expenses (log)	-0.003 (0.01)	0.005 (0.01)	-0.002 (0.01)	-0.002 (0.02)	-0.002 (0.02)	0.003 (0.02)	0.004 (0.01)	-0.002 (0.01)	-0.002 (0.01)	0.018 (0.02)	-0.029 (0.02)	0.011 (0.02)
Nuclear family	-0.000 (0.01)	-0.031* (0.02)	0.031 (0.02)	0.016 (0.02)	-0.038* (0.02)	0.023 (0.02)	-0.019 (0.02)	0.021 (0.01)	-0.002 (0.02)	-0.018 (0.02)	0.033 (0.02)	-0.015 (0.02)
Formal credit	-0.016 (0.02)	-0.004 (0.02)	0.020 (0.01)	-0.001 (0.02)	-0.007 (0.02)	0.008 (0.02)	-0.010 (0.02)	0.036** (0.02)	-0.025 (0.02)	-0.001 (0.02)	0.020 (0.02)	-0.019 (0.02)
Get help	-0.020 (0.02)	0.015 (0.02)	0.004 (0.02)	-0.011 (0.02)	0.018 (0.02)	-0.008 (0.02)	0.013 (0.02)	-0.012 (0.02)	-0.001 (0.02)	0.021 (0.02)	-0.021 (0.02)	-0.000 (0.02)
Send help	-0.001 (0.02)	0.018 (0.02)	-0.017 (0.02)	-0.011 (0.02)	0.019 (0.02)	-0.008 (0.02)	0.020 (0.03)	0.011 (0.02)	-0.031 (0.02)	0.016 (0.03)	0.025 (0.02)	-0.041* (0.02)
Social program beneficiary	0.018 (0.01)	0.003 (0.01)	-0.020 (0.01)	0.029 (0.03)	-0.021 (0.03)	-0.008 (0.03)	0.003 (0.02)	0.000 (0.02)	-0.003 (0.01)	-0.008 (0.03)	0.029 (0.02)	-0.021 (0.03)
FEA	0.009 (0.02)	0.012 (0.02)	-0.021 (0.02)	-0.021 (0.04)	0.033 (0.04)	-0.012 (0.03)	0.006 (0.02)	-0.012 (0.02)	0.005 (0.02)	0.008 (0.03)	-0.048* (0.03)	0.041 (0.03)
Stratum 1	0.001 (0.02)	0.003 (0.02)	-0.004 (0.02)	-0.007 (0.02)	0.021 (0.03)	-0.014 (0.03)						
Stratum 2	-0.002 (0.01)	0.007 (0.02)	-0.005 (0.01)	-0.002 (0.02)	0.015 (0.02)	-0.014 (0.02)						
Wealth (rural): quintile 1							0.025 (0.02)	-0.016 (0.02)	-0.009 (0.02)	0.018 (0.04)	-0.035 (0.04)	0.017 (0.04)
Wealth (rural): quintile 2							0.005 (0.02)	0.006 (0.02)	-0.011 (0.02)	0.005 (0.04)	-0.016 (0.03)	0.011 (0.04)
Wealth (rural): quintile 3							0.012 (0.02)	-0.029 (0.02)	0.017 (0.02)	0.011 (0.03)	-0.038 (0.03)	0.027 (0.04)
Wealth (rural): quintile 4							-0.014 (0.02)	0.017 (0.02)	-0.003 (0.02)	-0.006 (0.03)	-0.006 (0.03)	0.011 (0.03)
No sewage							0.020 (0.02)	-0.014 (0.02)	-0.006 (0.02)	0.024 (0.02)	-0.014 (0.02)	-0.010 (0.02)

*Notes: The table reports marginal effects from multinomial probit models, with standard errors in parentheses. All regressions include region fixed effects. Treatment refers to respondents who were presented with the list that included a sensitive item, Control 1 received the list without the sensitive item followed by the direct question, and Control 2 was asked the direct question. For variable definitions, see Appendix Table A-1. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.*



**Table A-3: Balance on covariates at the follow-up (2013): vote-buying list experiment**

Variables	Urban sample						Rural sample					
	Bivariate			Multivariate			Bivariate			Multivariate		
	Treatment	Control 1	Control 2	Treatment	Control 1	Control 2	Treatment	Control 1	Control 2	Treatment	Control 1	Control 2
Age	0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	0.000 (0.00)	0.000 (0.00)	-0.001 (0.00)	-0.000 (0.00)	-0.001 (0.00)	0.001 (0.00)	0.000 (0.00)	-0.001** (0.00)	0.001 (0.00)
Male household head	0.002 (0.01)	-0.006 (0.01)	0.004 (0.01)	0.036* (0.02)	-0.014 (0.02)	-0.021 (0.02)	-0.010 (0.02)	0.003 (0.02)	0.007 (0.02)	-0.007 (0.02)	0.022 (0.02)	-0.015 (0.02)
Education	-0.000 (0.00)	0.001 (0.00)	-0.000 (0.00)	0.002 (0.00)	0.001 (0.00)	-0.003 (0.00)	-0.001 (0.00)	0.001 (0.00)	-0.000 (0.00)	-0.001 (0.00)	-0.001 (0.00)	0.002 (0.00)
Employed household head	-0.028* (0.02)	0.041** (0.02)	-0.013 (0.02)	-0.028 (0.02)	0.047** (0.02)	-0.019 (0.02)	0.029 (0.02)	-0.053*** (0.02)	0.024 (0.02)	0.037* (0.02)	-0.072*** (0.02)	0.035* (0.02)
Savings	-0.019 (0.03)	-0.004 (0.02)	0.023 (0.02)	-0.013 (0.03)	-0.009 (0.02)	0.023 (0.02)	0.006 (0.02)	0.002 (0.02)	-0.008 (0.02)	0.011 (0.02)	-0.004 (0.02)	-0.008 (0.02)
Not in organization	-0.026* (0.02)	0.033** (0.02)	-0.007 (0.02)	-0.017 (0.02)	0.032* (0.02)	-0.014 (0.02)	0.007 (0.02)	-0.012 (0.02)	0.005 (0.02)	0.008 (0.02)	-0.012 (0.02)	0.003 (0.02)
Social security	-0.036 (0.03)	0.074** (0.04)	-0.038 (0.03)	-0.037 (0.03)	0.069* (0.04)	-0.032 (0.03)	-0.032 (0.03)	0.035 (0.03)	-0.002 (0.03)	-0.033 (0.04)	0.024 (0.03)	0.009 (0.04)
Contributes to social security	-0.032** (0.02)	-0.003 (0.01)	0.035** (0.01)	-0.036** (0.02)	-0.002 (0.02)	0.038** (0.02)	0.007 (0.02)	-0.018 (0.02)	0.011 (0.02)	-0.000 (0.02)	-0.016 (0.02)	0.016 (0.02)
Home with spouse	0.015 (0.01)	0.003 (0.01)	-0.017 (0.01)	0.031 (0.02)	-0.004 (0.02)	-0.027 (0.02)	0.024 (0.02)	-0.011 (0.02)	-0.013 (0.02)	0.035 (0.02)	-0.002 (0.02)	-0.033 (0.02)
Wealth	-0.013* (0.01)	0.009 (0.01)	0.003 (0.01)	-0.008 (0.01)	0.015 (0.01)	-0.006 (0.01)	-0.005 (0.01)	0.010 (0.01)	-0.005 (0.01)	-0.002 (0.02)	-0.003 (0.02)	0.005 (0.02)
People in household	0.001 (0.00)	-0.003 (0.00)	0.003 (0.00)	0.000 (0.00)	-0.003 (0.01)	0.003 (0.01)	0.001 (0.00)	0.004 (0.00)	-0.005 (0.00)	-0.000 (0.00)	0.006 (0.00)	-0.006 (0.00)
Overcrowded	-0.035** (0.02)	0.043** (0.02)	-0.008 (0.02)	-0.028 (0.02)	0.038 (0.03)	-0.010 (0.02)	-0.017 (0.02)	0.012 (0.02)	0.004 (0.02)	-0.012 (0.02)	0.023 (0.02)	-0.011 (0.02)
Homeowner	0.001 (0.01)	-0.005 (0.01)	0.003 (0.01)	0.006 (0.02)	-0.011 (0.02)	0.005 (0.01)	-0.022 (0.01)	0.005 (0.01)	0.018 (0.01)	-0.022 (0.02)	0.004 (0.02)	0.017 (0.02)
No debts	0.031** (0.01)	-0.011 (0.01)	-0.020 (0.01)	0.051* (0.03)	-0.033 (0.03)	-0.017 (0.02)	-0.000 (0.01)	0.011 (0.01)	-0.011 (0.01)	-0.029 (0.03)	0.020 (0.02)	0.009 (0.03)
Shock	0.039*** (0.01)	-0.026* (0.01)	-0.013 (0.01)	0.030** (0.01)	-0.015 (0.01)	-0.015 (0.01)	0.021 (0.02)	0.007 (0.02)	-0.029* (0.02)	0.021 (0.02)	0.005 (0.02)	-0.026 (0.02)
HH expenses (log)	-0.008 (0.01)	-0.004 (0.01)	0.012 (0.01)	0.007 (0.02)	-0.028 (0.02)	0.021 (0.02)	0.010 (0.01)	0.010 (0.01)	-0.020* (0.01)	-0.018 (0.01)	0.023 (0.02)	-0.005 (0.02)
HH food expenses (log)	-0.010 (0.01)	0.004 (0.01)	0.006 (0.01)	-0.009 (0.02)	0.022 (0.02)	-0.014 (0.02)	0.022* (0.01)	0.001 (0.01)	-0.023** (0.01)	0.043** (0.01)	-0.025 (0.02)	-0.018 (0.02)
Nuclear family	0.005 (0.01)	0.002 (0.01)	-0.007 (0.01)	0.015 (0.02)	-0.004 (0.02)	-0.011 (0.02)	-0.006 (0.02)	-0.000 (0.02)	0.007 (0.02)	0.002 (0.02)	0.010 (0.02)	-0.012 (0.02)
Formal credit	0.014 (0.01)	0.001 (0.01)	-0.015 (0.01)	-0.025 (0.03)	0.031 (0.03)	-0.006 (0.03)	0.009 (0.02)	0.005 (0.01)	-0.013 (0.01)	0.036 (0.01)	-0.021 (0.03)	-0.014 (0.03)
Get help	-0.029** (0.01)	0.002 (0.01)	0.026* (0.01)	-0.018 (0.02)	-0.004 (0.02)	0.022 (0.02)	0.004 (0.02)	0.003 (0.02)	-0.007 (0.02)	0.008 (0.02)	0.002 (0.02)	-0.010 (0.02)
Send help	-0.022 (0.02)	0.028* (0.01)	-0.006 (0.02)	-0.021 (0.02)	0.028* (0.02)	-0.007 (0.02)	0.000 (0.02)	-0.004 (0.02)	0.004 (0.02)	0.001 (0.02)	-0.004 (0.02)	0.003 (0.02)
Social program beneficiary	0.021 (0.01)	-0.011 (0.01)	-0.010 (0.02)	0.017 (0.02)	-0.009 (0.02)	-0.008 (0.02)	0.008 (0.02)	0.006 (0.01)	-0.014 (0.01)	0.000 (0.02)	0.012 (0.02)	-0.012 (0.02)
FEA	0.014 (0.02)	-0.003 (0.02)	-0.011 (0.03)	-0.008 (0.03)	0.012 (0.03)	-0.005 (0.02)	0.013 (0.02)	-0.007 (0.02)	-0.006 (0.01)	0.012 (0.02)	-0.026 (0.02)	0.014 (0.02)
Stratum 1	-0.010 (0.02)	0.009 (0.02)	0.001 (0.01)	-0.025 (0.02)	0.026 (0.02)	-0.001 (0.02)						
Stratum 2	0.019 (0.02)	-0.004 (0.02)	-0.015 (0.01)	0.012 (0.02)	0.005 (0.02)	-0.017 (0.02)						
Wealth (rural): quintile 1							-0.008 (0.02)	-0.003 (0.02)	0.011 (0.02)	0.005 (0.05)	-0.027 (0.05)	0.022 (0.05)
Wealth (rural): quintile 2							0.034* (0.02)	-0.045** (0.02)	0.010 (0.02)	0.035 (0.05)	-0.056 (0.04)	0.021 (0.04)
Wealth (rural): quintile 3							0.010 (0.02)	0.009 (0.02)	-0.019 (0.02)	0.015 (0.04)	-0.013 (0.03)	-0.002 (0.04)
Wealth (rural): quintile 4							-0.031* (0.02)	0.020 (0.02)	0.011 (0.02)	-0.019 (0.03)	-0.001 (0.03)	0.021 (0.03)
No sewage							0.020 (0.02)	-0.005 (0.02)	-0.015 (0.02)	0.024 (0.02)	-0.008 (0.02)	-0.016 (0.02)

*Notes: The table reports marginal effects from multinomial probit models, with standard errors in parentheses. All regressions include region fixed effects. Treatment refers to respondents who were presented with the list that included a sensitive item, Control 1 received the list without the sensitive item followed by the direct question, and Control 2 was asked the direct question. For variable definitions, see Appendix Table A-1. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.*

### A.3 Testing the no design and no liar assumptions in list experiments

In this Appendix we test for the ‘no design’ and ‘no liar’ assumptions in our list experiments, following Blair and Imai (2012). The test for the former compares the predicted average difference in answers to control items under treatment vs. control. With  $Y_i(0)$ ,  $Z_{i,J+1}^*$ ,  $Y_i$  and  $T_i$  specified as above, let  $\pi_{yz} = \Pr(Y_i(0), Z_{i,J+1}^* = z)$  represent the proportion of the population in each type  $(Y_i(0), Z_{i,J+1}^*)$ . If there are no design effects, these proportions can be computed for all  $y = 0, \dots, J$  as follows:

$$\begin{aligned}\pi_{y1} &= \Pr(Y_i \leq y | T_i = 0) - \Pr(Y_i \leq y | T_i = 1), \\ \pi_{y0} &= \Pr(Y_i \leq y | T_i = 1) - \Pr(Y_i \leq y - 1 | T_i = 0).\end{aligned}$$

Proportions  $\pi_{y1}$  and  $\pi_{y0}$  always take positive values. But with design effects, estimated proportions can be negative (for example, see Table 5 in Blair and Imai (2012)). To test for design effects, one can therefore evaluate whether the proportion of the population in each type ( $\pi_{yz}$ ) is jointly nonnegative.<sup>39</sup> Panel A in Table A-4 shows that no single estimated proportion is negative for either experiment, so the test suggests there is no evidence to reject the null hypothesis of no design effects.

To test the ‘no liar’ assumption, we can evaluate the two most common sources of untruthful answers: ceiling and floor effects. These occur when the respondent engages in either none or all of the behaviors, and thus feels exposed if he or she answers truthfully. In Table 1, the bulk of the answers in the treated lists (93.9%) are larger than zero and smaller than the maximum (five) number of items people can list. This reflects that, since the original instrument design, we included option items that are likely to be negatively correlated with each other, as well as at least one very frequent behavior.

We also test for floor and ceiling effects more formally by estimating

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<sup>39</sup>This test, however, has limitations: there can be design effects with positive  $\pi_{y1}$  and  $\pi_{y0}$ . Also, a higher probability of positive answers to the sensitive item reduces the likelihood of rejecting the null of no design effects.

the model under the no liar assumption, and comparing it to an alternative model allowing for floor and ceiling effects. Based on different information criteria, if the data supports the second model, there is evidence to reject the null of no floor or ceiling effects.<sup>40</sup> Panel B of Table A-4 reports the results. Regardless of the criterion used, Schwarz's BIC or Akaike's AIC, the preferred model includes no floor or ceiling effects, so this test fails to reject the null of no floor or ceiling effects. Furthermore, these results hold either with covariates (Columns 1 and 2) or when the basic set of covariates in Figure 2 are included.<sup>41</sup>

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<sup>40</sup>Since the model is identified under the no floor or ceiling effects assumption, we must make additional assumptions to estimate the alternative, allowing for these effects. To do so, we follow Blair and Imai (2012) and consider that respondents' truthful answers to the sensitive item are independent of their answers for control items, conditional upon the pretreatment covariates.

<sup>41</sup>We also find similar results using a different set of covariates.

**Table A-4: Testing assumptions in the list experiments**

	(1)	(2)	(3)	(4)
<b>Panel A: No design effects</b>				
Response value ( $y$ )	Estimated proportions with response $y$ to control items and...			
	...not following sensitive behavior ( $\hat{\pi}_{y0}$ )		...following sensitive behavior ( $\hat{\pi}_{y1}$ )	
	<i>Estimate</i>	<i>Std. Error</i>	<i>Estimate</i>	<i>Std. Error</i>
0	0.045	0.004	0.020	0.006
1	0.407	0.011	0.055	0.014
2	0.288	0.012	0.053	0.010
3	0.067	0.008	0.015	0.006
4	0.033	0.005	0.015	0.002
Total	0.841		0.159	
P-value	1			
<b>Panel B: No liar effects</b>				
	Information criterion			
	<i>BIC</i>	<i>AIC</i>	<i>BIC</i>	<i>AIC</i>
<i>Clientelism</i>				
No boundary	9873.01	9885.44	9863.88	10050.28
Ceiling	9875.50	9894.14	9894.20	10173.80
Floor	9896.91	9915.55	9897.23	10176.84
Ceiling-Floor	9899.40	9924.26	9927.55	10300.36

**Notes:** Panel A reports the estimated proportion of respondent types as described in each column title. The design effects test evaluates whether the population proportions are jointly non-negative. For each experiment, the Bonferroni-corrected P-value for the null of no design effects is reported. Panel B reports Schwarz's (BIC) and Akaike's (AIC) information criteria when the model is estimated without including boundaries (No boundary), including ceiling effects (Ceiling), including floor effects (Floor) and including both ceiling and floor effects (Ceiling-Floor). In this panel, the first two columns estimate the models without covariates, while the final two columns include the set of characteristics listed in Figure 2.

## A.4 Simple regression analysis

In the main text, we focus on the extreme bounds methodology to examine which variables are robustly correlated with clientelism. This section reports a simpler regression analysis, which produces similar conclusions.

Table A-5 runs linear regressions for clientelism on the same set of variables explored in the text. Odd columns, labeled “bivariate”, show the resulting coefficient for regressions including only one covariate at a time (in addition to region fixed effects, which are always included). Even columns show the coefficient for a multivariate regression, which simultaneously includes all variables listed in the table. As in our baseline analysis, variables are standardized to ease interpretation of the magnitude of the correlations. In the main text we also explored the role of a few interactions between correlates of interest. In Table A-6 we show the results of including such interaction terms in regressions for clientelism that include only region fixed effects and the relevant lower-order uninteracted terms (in the even, “bivariate” columns) as well as in regressions containing the full set of covariates in Table A-5. Again, there are few differences relative to the results using the extreme bounds methodology.

**Table A-5: Correlates of clientelism**  
Simple regression analysis

	(1)	(2)		(3)	(4)
<i>Variables</i>	<b>Bivariate</b>	<b>Multivariate</b>	<i>Variables</i>	<b>Bivariate</b>	<b>Multivariate</b>
Persuasion	0.117*** (0.0149)	0.108*** (0.0153)	Shock	0.0136 (0.0143)	0.00467 (0.0145)
Agree with bribery	0.0831*** (0.0150)	0.0730*** (0.0154)	Rural population	-0.0139 (0.0239)	-0.104*** (0.0224)
Other religion	-0.0471*** (0.00937)	-0.0441*** (0.0164)	Homicide rate	0.0159 (0.0183)	-0.00410 (0.0246)
Negative reciprocity	0.0659*** (0.0141)	0.0509*** (0.0140)	Neighbor cell phones	0.0122 (0.0143)	0.00222 (0.0146)
Party recall	0.0645*** (0.0125)	0.0576*** (0.0131)	Independent	0.0105 (0.0145)	0.00320 (0.0143)
Wealth	-0.0835*** (0.0203)	-0.0976*** (0.0272)	Neighbor loans	0.0127 (0.0186)	0.00882 (0.0204)
Positive reciprocity	0.0292*** (0.00958)	0.0244** (0.00975)	Right dominated	0.00733 (0.0136)	0.0151 (0.0149)
Right Ideology	0.0392*** (0.0143)	0.0215 (0.0146)	Left dominated	-0.0208 (0.0350)	0.0335 (0.0404)
Justice into own hands	0.0378*** (0.0129)	0.00617 (0.0138)	State presence	-0.00712 (0.0159)	-0.0124 (0.0177)
Age	-0.0272** (0.0129)	-0.0261* (0.0142)	Education	0.00405 (0.0131)	0.00976 (0.0151)
Secret ballot	-0.0281** (0.0135)	-0.0216 (0.0133)	Guerrillas	0.00403 (0.00974)	0.00666 (0.00951)
Use of violence	0.0334** (0.0136)	0.00281 (0.0141)	Own welfare	-0.00507 (0.0127)	-2.03e-05 (0.0138)
Non left-right contender	0.0319* (0.0163)	0.0492*** (0.0174)	Gov. against inequality	-0.00456 (0.0140)	-0.00622 (0.0135)
Party identity	0.0271* (0.0148)	-0.000311 (0.0158)	Government role	0.00531 (0.0147)	-0.000734 (0.0142)
Authorities violate law	0.0232* (0.0135)	-0.00652 (0.0145)	Popular vote	-0.00501 (0.0157)	-0.00337 (0.0153)
Win margin	-0.0211 (0.0154)	-0.0684** (0.0308)	Catholic	0.00249 (0.0128)	0.000538 (0.0344)
Polarization	-0.0235 (0.0199)	-0.105** (0.0510)	Lands	-0.00405 (0.0150)	-0.00145 (0.0156)
Evangelical/Pentecostal	0.0128 (0.0120)	0.0122 (0.0290)	Paramilitaries	0.00235 (0.0109)	-0.00784 (0.0123)
HH expenses	-0.0182 (0.0168)	-0.0113 (0.0179)	Left ideology	-0.00102 (0.0131)	-0.0109 (0.0143)
Pop. density	-0.0219 (0.0193)	-0.0292 (0.0249)	Woman	-0.00163 (0.0153)	0.0136 (0.0164)
Fractionalization	0.0170 (0.0169)	-0.109** (0.0553)	Gini of land properties	0.000559 (0.0162)	-0.00799 (0.0227)

*Notes:* Ordinary least squares regressions. The dependent variable of interest is a dummy indicating whether, when deciding who to vote for, the respondent has taken into account the benefits, gifts, or jobs that a candidate offered in exchange for the vote. Region fixed effects are always included, and standard errors are clustered at the community level. Odd columns, labeled “bivariate”, show the resulting coefficient for regressions including only one covariate at a time. Even columns show the coefficient for a multivariate regression, simultaneously including all variables listed in the table. For variable definitions, see Appendix Table A-1. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

**Table A-6: Clientelism: interaction terms**  
**Simple regression analysis**

<i>Variables</i>	(1)	(2)
	<b>Bivariate</b>	<b>Multivariate</b>
Right ideology × Right dominated	0.0416*** (0.0157)	0.0427** (0.0166)
Win margin × Negative rec.	-0.0216 (0.0146)	-0.0226 (0.0146)
Secret ballot × Positive rec.	-0.00980 (0.00955)	-0.0112 (0.00946)
Left ideology × Left dominated	0.00893 (0.0100)	0.00794 (0.00940)
Secret ballot × Negative rec.	-0.0117 (0.0145)	-0.0153 (0.0146)
Win margin × Positive rec.	-0.00332 (0.00984)	-0.00486 (0.00998)
Win margin × Wealth	0.00570 (0.0160)	0.00156 (0.0160)
Secret ballot × Rural population	0.00317 (0.0125)	0.00669 (0.0126)
Win margin × Secret ballot	0.00364 (0.0133)	0.00274 (0.0133)
Win margin × Rural pop.	0.00343 (0.0162)	-0.0137 (0.0186)
Secret ballot × Neighbor cell phones	-0.00172 (0.0141)	-0.00241 (0.0136)

*Notes: Ordinary least squares regressions. The dependent variable of interest is a dummy indicating whether, when deciding who to vote for, the respondent has taken into account the benefits, gifts, or jobs that a candidate offered in exchange for the vote. Standard errors are clustered at the community level. Region fixed effects are always included, and standard errors are clustered at the community level. Column 1 reports the coefficient of a “bivariate regression” containing only the region fixed effects, lower-order uninteracted terms, and the interaction of interest as regressors. Column 2 presents the results of a multivariate regression in which all variables in Table A-5. For variable definitions, see Appendix Table A-1. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.*

**Table A-7: Point estimates of the incidence and social desirability bias of clientelism across different covariates**

		(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Variables		Urban sample			Rural sample			Full sample		
		List	Direct	SDB	List	Direct	SDB	List	Direct	SDB
Age	(18,40]	0.172*** (0.049)	0.161*** (0.012)	0.012 (0.051)	0.155*** (0.055)	0.209*** (0.014)	-0.054 (0.057)	0.143*** (0.033)	0.182*** (0.009)	-0.039 (0.034)
	(40,55]	0.138*** (0.049)	0.138*** (0.010)	0.000 (0.050)	0.288*** (0.053)	0.238*** (0.013)	0.051 (0.054)	0.218*** (0.035)	0.186*** (0.008)	0.032 (0.036)
	>55	0.219*** (0.065)	0.155*** (0.015)	0.064 (0.066)	0.244*** (0.055)	0.210*** (0.015)	0.035 (0.057)	0.213*** (0.040)	0.185*** (0.011)	0.027 (0.041)
Education	Primary or less	0.227*** (0.054)	0.154*** (0.012)	0.073 (0.055)	0.247*** (0.038)	0.220*** (0.009)	0.027 (0.039)	0.218*** (0.030)	0.200*** (0.007)	0.019 (0.030)
	Secondary	0.112*** (0.041)	0.145*** (0.011)	-0.033 (0.042)	0.199*** (0.066)	0.218*** (0.017)	-0.019 (0.069)	0.161*** (0.037)	0.168*** (0.009)	-0.008 (0.038)
	College	0.187*** (0.066)	0.152*** (0.014)	0.035 (0.067)	0.172 (0.139)	0.260*** (0.049)	-0.088 (0.146)	0.158*** (0.052)	0.163*** (0.014)	-0.004 (0.054)
Gender	Men	0.161*** (0.047)	0.141*** (0.011)	0.020 (0.048)	0.209*** (0.045)	0.232*** (0.012)	-0.023 (0.047)	0.175*** (0.031)	0.191*** (0.008)	-0.016 (0.032)
	Women	0.172*** (0.037)	0.155*** (0.009)	0.018 (0.039)	0.256*** (0.042)	0.210*** (0.011)	0.046 (0.044)	0.203*** (0.028)	0.179*** (0.007)	0.024 (0.029)
Employment	No	0.147*** (0.051)	0.133*** (0.013)	0.014 (0.053)	0.303*** (0.059)	0.205*** (0.014)	0.098 (0.061)	0.217*** (0.040)	0.171*** (0.009)	0.046 (0.041)
	Yes	0.176*** (0.036)	0.156*** (0.008)	0.020 (0.037)	0.200*** (0.037)	0.229*** (0.010)	-0.029 (0.039)	0.179*** (0.024)	0.190*** (0.007)	-0.011 (0.025)
Frequency of voting	Low	0.138** (0.053)	0.116*** (0.012)	0.022 (0.054)	0.117 (0.075)	0.161*** (0.018)	-0.044 (0.077)	0.126*** (0.040)	0.132*** (0.010)	-0.006 (0.042)
	High	0.181*** (0.035)	0.162*** (0.008)	0.019 (0.036)	0.256*** (0.035)	0.232*** (0.009)	0.025 (0.036)	0.209*** (0.024)	0.199*** (0.006)	0.011 (0.025)
Vote for the same party	No	0.149*** (0.036)	0.147*** (0.008)	0.002 (0.037)	0.199*** (0.045)	0.219*** (0.011)	-0.020 (0.047)	0.161*** (0.027)	0.177*** (0.007)	-0.016 (0.027)
	Yes	0.213*** (0.051)	0.155*** (0.013)	0.058 (0.052)	0.272*** (0.044)	0.222*** (0.012)	0.050 (0.045)	0.238*** (0.034)	0.196*** (0.009)	0.042 (0.035)
Secret ballot	No	0.198*** (0.057)	0.165*** (0.013)	0.033 (0.059)	0.243*** (0.059)	0.257*** (0.017)	-0.014 (0.061)	0.192*** (0.038)	0.205*** (0.010)	-0.013 (0.039)
	Yes	0.153*** (0.037)	0.142*** (0.008)	0.011 (0.038)	0.231*** (0.036)	0.208*** (0.009)	0.023 (0.037)	0.191*** (0.026)	0.176*** (0.006)	0.015 (0.026)
Religion	None or other	0.104* (0.060)	0.162*** (0.016)	-0.058 (0.062)	0.264*** (0.084)	0.295*** (0.023)	-0.031 (0.087)	0.174*** (0.049)	0.216*** (0.013)	-0.042 (0.051)
	Catholic	0.187*** (0.034)	0.146*** (0.008)	0.040 (0.035)	0.231*** (0.034)	0.208*** (0.009)	0.023 (0.035)	0.196*** (0.023)	0.178*** (0.006)	0.018 (0.024)
Skin color	White	0.165*** (0.037)	0.129*** (0.008)	0.036 (0.038)	0.229*** (0.037)	0.191*** (0.010)	0.038 (0.038)	0.192*** (0.026)	0.158*** (0.006)	0.034 (0.027)
	Black	0.173*** (0.047)	0.192*** (0.013)	-0.019 (0.049)	0.243*** (0.055)	0.269*** (0.014)	-0.026 (0.056)	0.192*** (0.035)	0.233*** (0.010)	-0.040 (0.037)
Shock	No	0.140*** (0.041)	0.123*** (0.011)	0.017 (0.043)	0.213*** (0.057)	0.225*** (0.016)	-0.012 (0.059)	0.169*** (0.035)	0.168*** (0.009)	-0.002 (0.036)
	Yes	0.183*** (0.040)	0.164*** (0.009)	0.018 (0.041)	0.242*** (0.036)	0.219*** (0.009)	0.023 (0.038)	0.202*** (0.026)	0.193*** (0.007)	0.009 (0.027)
Wealth	Below median	0.186*** (0.043)	0.186*** (0.011)	-0.001 (0.044)	0.229*** (0.041)	0.252*** (0.011)	-0.023 (0.042)	0.221*** (0.029)	0.225*** (0.008)	-0.004 (0.030)
	Above median	0.152*** (0.042)	0.118*** (0.009)	0.034 (0.043)	0.242*** (0.047)	0.183*** (0.011)	0.059 (0.049)	0.156*** (0.030)	0.136*** (0.007)	0.020 (0.031)

*Notes:* For each area, rural, urban, and the full sample, the table shows the average incidence of clientelism using the list experiment (with methods as described in Section 2.1), the direct question (using a logit model), and the difference between the two or social desirability bias –SDB– (following Equation (3)). Standard errors are computed using Monte Carlo simulations, and estimations control for the set of variables listed in the table. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level. See Table A-1 for a description of all variables.